Reason-Rupe September 2012 Poll Methodology

Prepared by Princeton Survey Research Associates International for Reason Foundation

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SUMMARY

The Reason-Rupe poll obtained telephone interviews with a nationally representative sample of 1,006 adults living in the continental United States. Telephone interviews were conducted by landline (602) and cell phone (404, including 208 without a landline). The survey was conducted by Princeton Survey Research International. The interviews were done in English by Princeton Data Source, LLC from September 13 to 17, 2012. Statistical results are weighted to correct known demographic discrepancies. The margin of sampling error for the complete set of weighted data is ± 3.8 percentage points.

Details on the design, execution and analysis of the survey are discussed below.

DESIGN AND DATA COLLECTION PROCEDURES

Sample Design

A combination of landline and cellular random digit dial (RDD) samples was used to represent all adults in the continental United States who have access to either a landline or cellular telephone. Both samples were provided by Survey Sampling International, LLC (SSI) according to PSRAI specifications.

Numbers for the landline sample were drawn with equal probabilities from active blocks (area code + exchange + two-digit block number) that contained three or more residential directory listings. The cellular sample was not list-assisted, but was drawn through a systematic sampling from dedicated wireless 100-blocks and shared service 100-blocks with no directory-listed landline numbers.

Contact Procedures

Interviews were conducted from September 13 to 17, 2012. As many as seven attempts were made to contact every sampled telephone number. Sample was released for interviewing in replicates, which are representative subsamples of the larger sample. Using replicates to control the release of sample ensures that complete call procedures are followed for the entire sample. Calls were staggered over times of day and days of the week to maximize the chance of making contact with potential respondents. Each phone number received at least one daytime call when necessary.

For the landline sample, interviewers asked to speak with the youngest adult male or female currently at home based on a random rotation. If no male/female was available, interviewers asked to speak with the youngest adult of the other gender. This systematic respondent selection technique has been shown to produce samples that closely mirror the population in terms of age and gender when combined with cell interviewing.

For the cellular sample, interviews were conducted with the person who answered the phone. Interviewers verified that the person was an adult and in a safe place before administering the survey.

WEIGHTING AND ANALYSIS

Weighting is generally used in survey analysis to compensate for sample designs and patterns of non-response that might bias results. The sample was weighted to match national adult general population parameters. A two-stage weighting procedure was used to weight this dual-frame sample.

The first stage of weighting corrected for different probabilities of selection associated with the number of adults in each household and each respondent's telephone usage patterns.¹ This weighting also adjusts for the overlapping landline and cell sample frames and the relative sizes of each frame and each sample.

¹ i.e., whether respondents have only a landline telephone, only a cell phone, or both kinds of telephone.

This first-stage weight for the ith case can be expressed as:

$$WT_{i} = \frac{1}{\left(\frac{S_{LL}}{S_{CP}} \times \frac{1}{AD_{i}}\right)} \text{ if respondent has no cell phone}$$
$$WT_{i} = \frac{1}{\left(\frac{S_{LL}}{S_{CP}} \times \frac{1}{AD_{i}}\right) + R} \text{ if respondent has both kinds of phone}$$
$$WT_{i} = \frac{1}{R} \text{ if respondent has no lan line phone}$$

Where S_{LL} = size of the landline sample

 S_{CP} = size of the cell phone sample

 $AD_i = Number of adults in the household$

R = Estimated ratio of the land line sample frame to the cell phone sample frame

The equations can be simplified by plugging in the values for $S_{LL} = 602$ and $S_{CP} = 404$. Additionally, we will estimate of the ratio of the size of landline sample frame to the cell phone sample frame R = 0.70.

The second stage of weighting balanced sample demographics to population parameters. The sample is balanced to match national population parameters for sex, age, education, race, Hispanic origin, region (U.S. Census definitions), population density, number of adults in household, telephone usage, employment status and party identification. The basic weighting parameters came from a special analysis of the Census Bureau's 2011 Annual Social and Economic Supplement (ASEC) that included all households in the continental United States. The population density parameter was derived from Census data. The telephone usage parameter came from an analysis of the July-December 2011 National Health Interview Survey.² The party identification parameter was derived from recent PSRAI Omnibus survey data.

Weighting was accomplished using Sample Balancing, a special iterative sample weighting program that simultaneously balances the distributions of all variables using a statistical technique called the *Deming Algorithm*. Weights were trimmed to prevent individual interviews from having too much influence on the final results. The use of these weights in statistical analysis ensures that the demographic characteristics of the sample closely approximate the demographic characteristics of the national population. Table 1 compares weighted and unweighted sample distributions to population parameters.

² Blumberg SJ, Luke JV. Wireless substitution: Early release of estimates from the National Health Interview Survey, July-December, 2011. National Center for Health Statistics. Jul 2012.

_	Parameter	Unweighted	Weighted
Gender			
Male	48.6	47.3	48.8
Female	51.4	52.7	51.2
Age			
18-24	12.8	11.0	12.7
25-34	18.0	10.0	17.1
35-44	17.2 19.0	11.8 19.1	16.9 19.8
45-54			
55-64	16.0	22.9	16.5
65+	17.0	25.2	17.1
Education (changed)			
Less than HS Graduate	13.3	5.2	10.8
HS Graduate	30.4	26.4	31.1
Some College/Assoc Degree	28.5	29.9	28.6
College Graduate	27.8	38.5	29.5
Race/Ethnicity			
White/not Hispanic	68.0	72.5	68.3
Black/not Hispanic	11.6	12.8	11.7
Hispanic	14.1	8.9	13.3
Other/not Hispanic	6.3	5.8	6.7
Region			
Northeast	18.5	13.1	17.8
Midwest	21.8	22.2	21.6
South	37.0	37.6	37.4
West	22.7	27.1	23.2
County Pop. Density			
1 - Lowest	20.1	24.8	19.7
2	20.0	22.1	20.5
3	20.1	21.7	20.4
4	20.2	17.5	19.7
5 - Highest	19.6	14.0	19.8
Household Phone Use			
LLO	7.0	6.7	7.1
Dual	57.8	72.7	59.4
СРО	35.2	20.7	33.5

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able 1: Sample Demographics (continued)				
	Parameter	<u>Unweighted</u>	Weighted	
<u># of adults in HH</u>				
One	17.0	26.9	17.8	
Two	52.9	48.0	53.0	
Three +	30.1	25.0	29.2	
Employment Status				
Employed - public	9.0	14.7	9.3	
Employed - private	44.0	29.5	42.8	
Self-employed	6.3	7.6	6.5	
Not employed	40.7	48.1	41.3	
Party ID				
Republican	23.7	26.2	23.8	
Democrat	31.8	37.4	32.6	
Independent	35.5	30.5	35.2	
None/Other	9.0	5.9	8.5	

Table 1: Sample Demographics (continued)

Effects of Sample Design on Statistical Inference

Post-data collection statistical adjustments require analysis procedures that reflect departures from simple random sampling. PSRAI calculates the effects of these design features so that an appropriate adjustment can be incorporated into tests of statistical significance when using these data. The so-called "design effect" or *deff* represents the loss in statistical efficiency that results from a disproportionate sample design and systematic non-response. The total sample design effect for this survey is 1.54.

PSRAI calculates the composite design effect for a sample of size n, with each case having a weight, w_i as:

$$deff = \frac{n \sum_{i=1}^{n} w_i^2}{\left(\sum_{i=1}^{n} w_i\right)^2} \qquad formula \ label{eq:deff}$$

In a wide range of situations, the adjusted *standard error* of a statistic should be calculated by multiplying the usual formula by the square root of the design effect (\sqrt{deff}). Thus, the formula for computing the 95% confidence interval around a percentage is:

$$\hat{p} \pm \left(\sqrt{deff} \times 1.96\sqrt{\frac{\hat{p}(1-\hat{p})}{n}}\right)$$
 formula 2

where \hat{p} is the sample estimate and *n* is the unweighted number of sample cases in the group being considered.

The survey's *margin of error* is the largest 95% confidence interval for any estimated proportion based on the total sample— the one around 50%. For example, the margin of error for the entire sample is ±3.8 percentage points. This means that in 95 out every 100 samples drawn using the same methodology, estimated proportions based on the entire sample will be no more than 3.8 percentage points away from their true values in the population. It is important to remember that sampling fluctuations are only one possible source of error in a survey estimate. Other sources, such as respondent selection bias, questionnaire wording and reporting inaccuracy, may contribute additional error of greater or lesser magnitude.

RESPONSE RATE

Table 2 report the disposition of all sampled telephone numbers ever dialed from the original telephone number samples. The response rate estimates the fraction of all eligible sample that was ultimately interviewed. At PSRAI it is calculated by taking the product of three component rates:³

- Contact rate the proportion of working numbers where a request for interview was made⁴
- Cooperation rate the proportion of contacted numbers where a consent for interview was at least initially obtained, versus those refused
- Completion rate the proportion of initially cooperating and eligible interviews that were completed

Thus the response rate for the land line samples was 9 percent. The response rate for the cellular samples was 7 percent.

³ PSRAI's disposition codes and reporting are consistent with the American Association for Public Opinion Research standards.

⁴ PSRAI assumes that 75 percent of cases that result in a constant disposition of "No answer" or "Busy" are actually not working numbers.

Table 2:Sample Disposition

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Landline	Cell			
29,616	23,895	Total Numbers Dialed		
712	291	Non-residential		
983	159	Computer/Fax		
6		Cell phone		
14,301	6,965	Other not working		
4,631	860	Additional projected not working		
8,984	15,621	Working numbers		
30.3%	65.4%	Working Rate		
1,544	287	No Answer / Busy		
2,083	8,603	Voice Mail		
144	58	Other Non-Contact		
5,213	6,673	Contacted numbers		
58.0%	42.7%	Contact Rate		
873	1,788	Callback		
3,538	3,767	Refusal		
802	1,118	Cooperating numbers		
15.4%	16.8%	Cooperation Rate		
183	360	Language Barrier		
	350	Child's cell phone		
619	408	Eligible numbers		
77.2%	36.5%	Eligibility Rate		
17	4	Break-off		
602	404	Completes		
97.3%	99.0%	Completion Rate		
8.7%	7.1%	Response Rate		